

CURVE FITTING AND DATA REGRESSION

Numerical methods in chemical
engineering

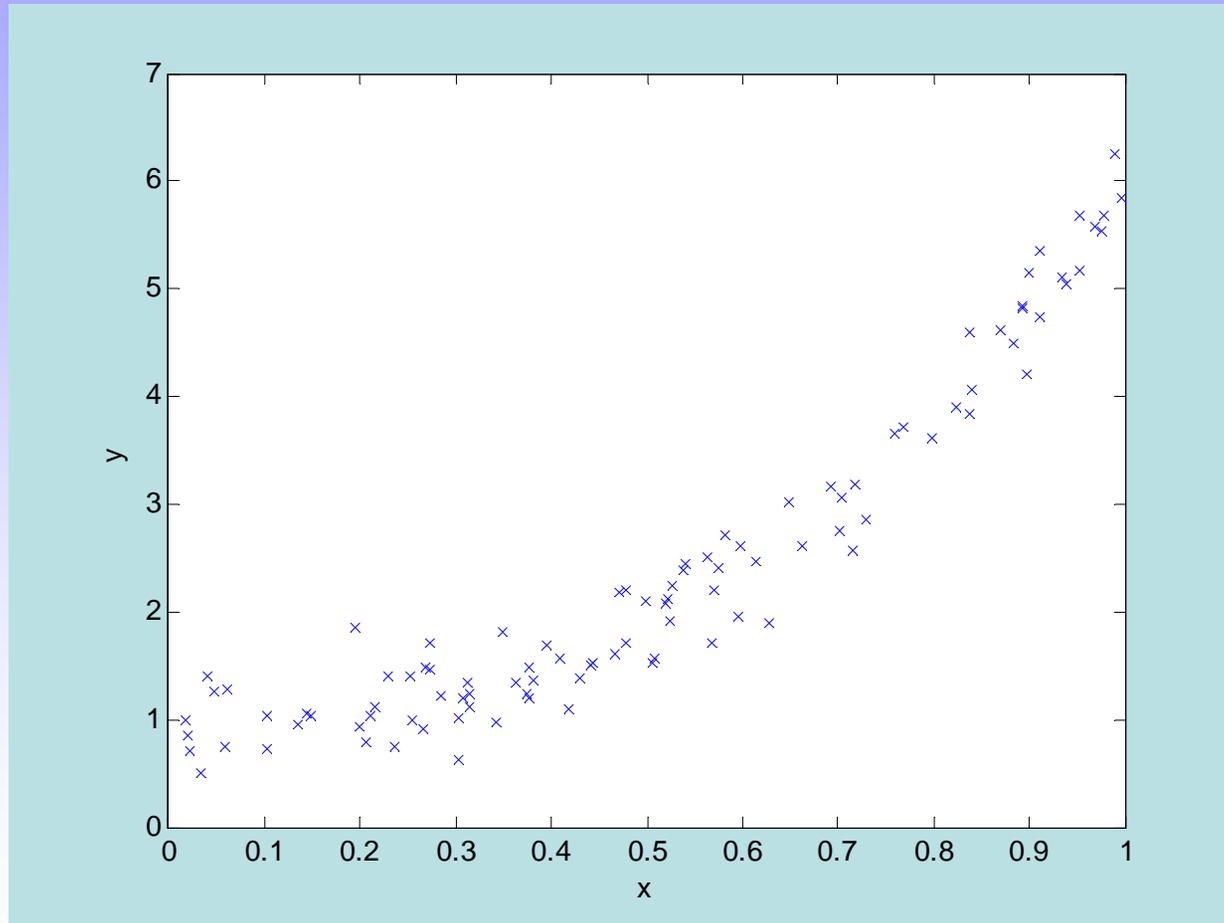
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OVERVIEW

- We are going to fit measurements to models today.
- You will also learn what R^2 actually means

FITTING MODELS TO DATA

y is the
measured
variable



x is the
controlled
variable

HOW TO FIT A MODEL TO THE DATA

- We would like to fit the following model to the data:

$$\hat{y} = a_1 + a_2x + a_3x^2 + a_4x^3 \quad (9-1)$$

- First step: If we have N data points, we could write the model as the product of a matrix and a vector:

$$\begin{bmatrix} \hat{y}_1 \\ \hat{y}_2 \\ \hat{y}_3 \\ \vdots \\ \hat{y}_N \end{bmatrix} = \begin{bmatrix} 1 & x_1 & x_1^2 & x_1^3 \\ 1 & x_2 & x_2^2 & x_2^3 \\ 1 & x_3 & x_3^2 & x_3^3 \\ \vdots & \vdots & \vdots & \vdots \\ 1 & x_N & x_N^2 & x_N^3 \end{bmatrix} \begin{bmatrix} a_1 \\ a_2 \\ a_3 \\ a_4 \end{bmatrix} \quad (9-2)$$

$$\hat{y} = Xa \quad (9-3)$$

X is called the design matrix and **a** are the fit parameters.

RESIDUALS

- Second step: work out the residuals for each data point:

$$d_i = (y_i - \hat{y}_i) \quad (9-4)$$

- Third step: Work out the sum of squares of the residuals:

$$\sum_i (d_i)^2 = \sum_i (y_i - \hat{y}_i)^2 \quad (9-5)$$

$$\sum_i (d_i)^2 = d \bullet d = d^T \times d = (y_i - \hat{y}_i)^T (y_i - \hat{y}_i) \quad (9-6)$$

MINIMIZING THE SUM OF SQUARES

- Choose the parameter vector such that the sum of squares of the residuals is minimized; the partial derivative with respect to each parameter should be set to zero:

$$\frac{\partial}{\partial a_j} [(y - (Xa))^T (y - Xa)] = 0 \quad (9-7)$$

MINIMIZING THE SUM OF SQUARES

$$\frac{\partial}{\partial a_j} [(y^T - (Xa)^T)(y - Xa)] = 0 \Leftrightarrow$$

$$\frac{\partial}{\partial a_j} [(y^T - X^T a^T)(y - Xa)] = 0 \Leftrightarrow$$

$$(y^T - X^T a^T)X \frac{\partial}{\partial a_j} [(a)] + \frac{\partial}{\partial a_j} [(a)^T] X^T (y - Xa) = 0 \Leftrightarrow \quad (9-8)$$

$$(y^T - X^T a^T)X e_j + e_j^T X^T (y - Xa) = 0$$

$$(y^T - (Xa)^T)X e_j + e_j^T X^T (y - Xa) =$$

$$(y - (Xa))^T X e_j + e_j^T X^T (y - Xa) = 0 \Leftrightarrow$$

$$(X e_j)^T (y - Xa) + e_j^T X^T (y - Xa) = 0 \Leftrightarrow$$

$$e_j^T X^T (y - Xa) + e_j^T X^T (y - Xa) = 0 \Leftrightarrow$$

$$e_j^T (X^T (y - Xa) + X^T (y - Xa)) = 0 \Leftrightarrow$$

$$2X^T (y - Xa) = 0 \Leftrightarrow$$

$$X^T y = X^T X a$$

$$a = (X^T X)^{-1} X^T y \quad (9-9)$$

USING MATLAB FOR LLSQ

- If we have the same number of data points as fit parameters, we can solve the

system:

$$\begin{bmatrix} y_1 \\ y_2 \\ y_3 \\ y_4 \end{bmatrix} = \begin{bmatrix} 1 & x_1 & x_1^2 & x_1^3 \\ 1 & x_2 & x_2^2 & x_2^3 \\ 1 & x_3 & x_3^2 & x_3^3 \\ 1 & x_4 & x_4^2 & x_4^3 \end{bmatrix} \begin{bmatrix} a_1 \\ a_2 \\ a_3 \\ a_4 \end{bmatrix} \quad (9-10)$$

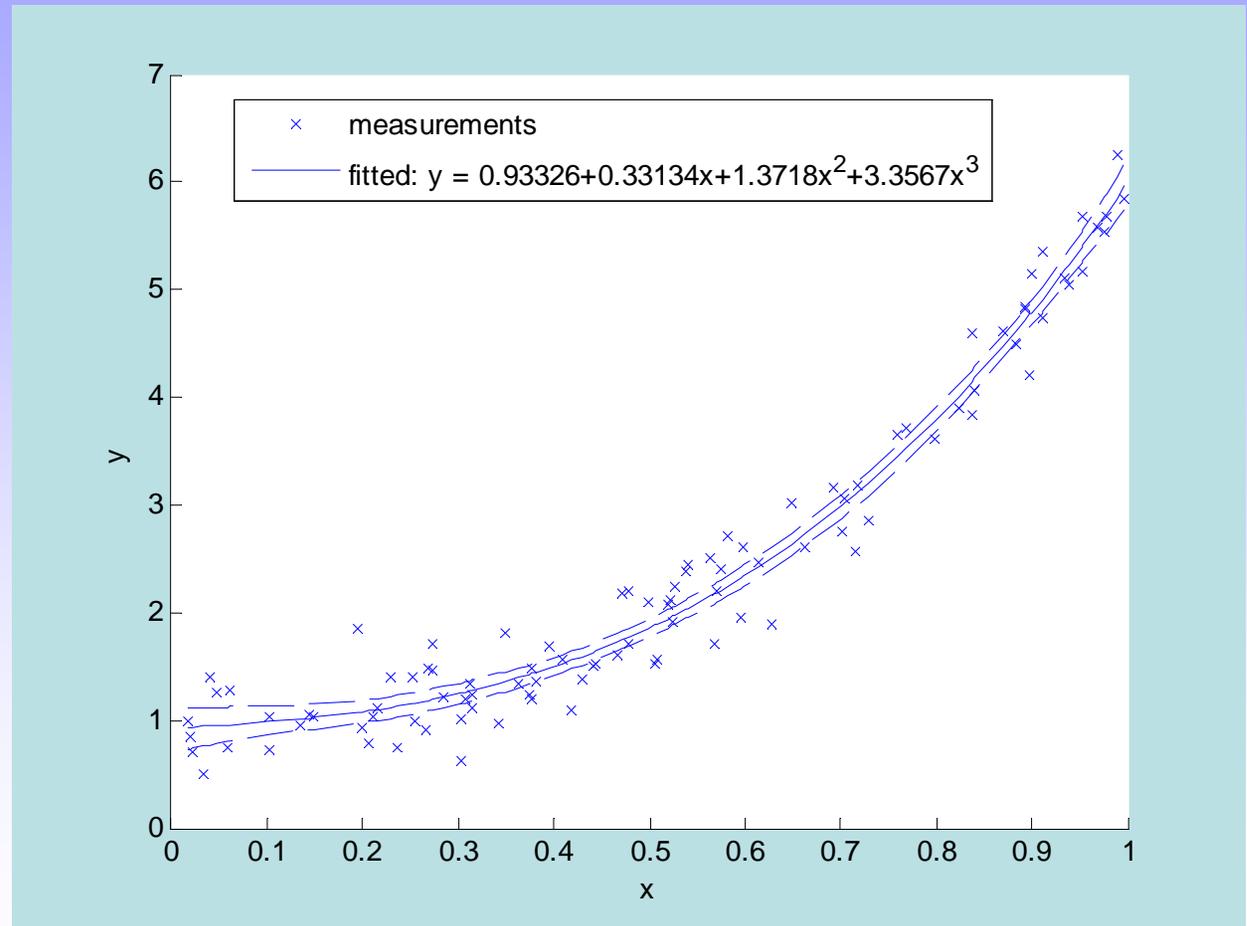
- As $a = X \backslash y$

USING MATLAB FPR LLSQ

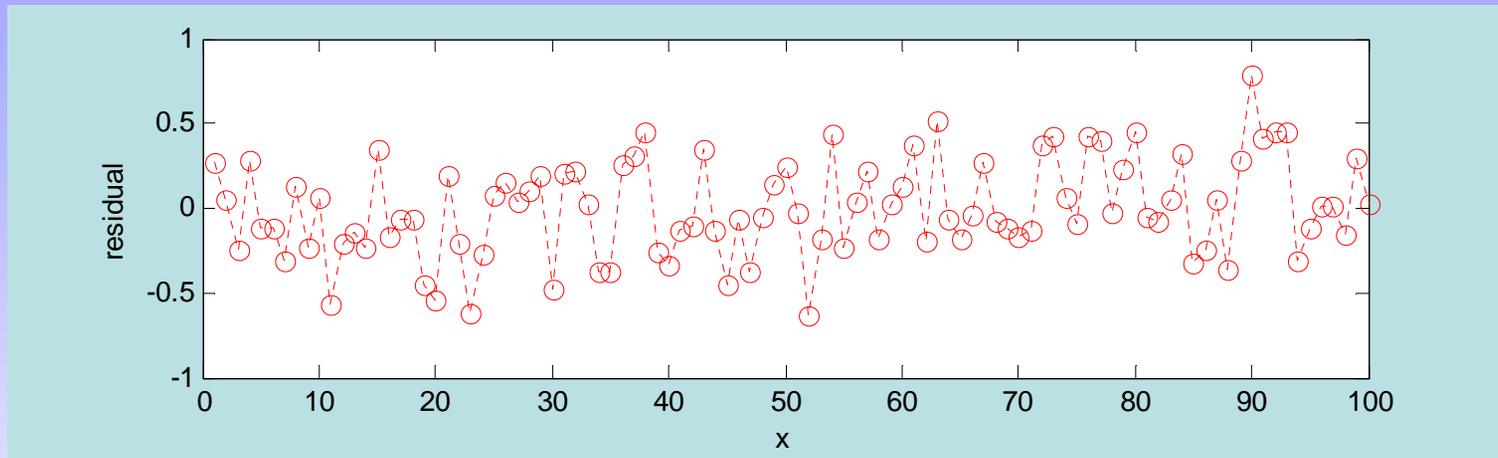
- If there are more data points ($N > 4$), we can write an analogue, but maybe a consistent solution does not exist (the system is over specified).
- However, matlab will find values for the vector a which minimize $\|y - aX\|^2$, so i.e. a least squares fit.

FIT TO OUR PROBLEM

```
N=length(x);  
X(:,1) = ones(N,1);  
X(:,2) = x;  
X(:,3) = x.^2;  
X(:,4) = x.^3;  
  
A =X\y;
```



HOW GOOD IS THE MODEL?



- For a model to make sense the data points should be scattered randomly around the model predictions, the mean of the residuals d should be zero.

$$d_i = (y_i - \hat{y}_i)$$

- It's always good to check if the residuals are not correlated with the measured values, if that is the case, it can indicate that your model is wrong.

REGRESSION COEFFICIENTS

- Variance measured in the data (y) is:

$$\sigma_y^2 = \frac{1}{N} \sum_i (y_i - \bar{y})^2 \quad (9-11)$$

- Variance of the residuals is:

$$\sigma_{error}^2 = \frac{1}{N} \sum_i (d_i)^2 \quad (9-12)$$

- Variance in the model is:

$$\sigma_{model}^2 = \frac{1}{N} \sum_i (\hat{y}_i - \bar{\hat{y}})^2 \quad (9-13)$$

REGRESSION COEFFICIENTS

- Given that the error is uncorrelated we can state that:

$$\sigma_y^2 = \sigma_{error}^2 + \sigma_{model}^2 \quad (9-14)$$

$$R^2 = \frac{\sigma_{model}^2}{\sigma_y^2} = 1 - \frac{\sigma_{error}^2}{\sigma_y^2} \quad (9-15)$$

$$R^2 = 1 - \frac{SSE}{SST} \quad (9-16)$$

SSE: Sum of errors squared
SSR: Sum of squares (data)
SST: Total sum of squares (model)

STATISTICAL ANALYSIS

- An uncorrelated error (mean will be zero) → SSE, SST and SSR will have χ^2 -distributions and the ratios will have an F -distribution. If SSR/SSE is large, the model is good!
- There is a change that the model is rubbish, but that SSR/SSE will yield a good value, Analysis of Variance (ANOVA) will be a good tool to calculate the probability of such a thing happening!

BACK TO THE EXAMPLE

- Stats:
 - $N = 100$
 - $SSE = 8.1031$
 - $SST = 232.5490$
 - $SSR = 224.4459$
 - $R^2 = 0.9652$

Source	Deg. Of freedom	Sum of squares	F-value
Regression	$K = 4$	$SSR = 224.44$	$F = 657.84$
Residual	$N-K-1 = 95$	$SSE = 8.103$	
Total	99	$SST = 232.55$	

$F > 657$ means: **very unlikely!!!**

$$F = (SSR/4)/(SSE/95)$$

CONFIDENCE LIMITS FOR THE FIT PARAMETERS

- Using the t-distribution, the confidence limits for the fit parameters can be set as:

$$a_j - t \frac{\sigma_{error}^2}{\nu} [(X^T X)^{-1}]_{j,j} < a_j < a_j + t \frac{\sigma_{error}^2}{\nu} [(X^T X)^{-1}]_{j,j} \quad (9-17)$$

Must be looked up from
a statistical table

Degree of freedom

Is the j-th diagonal
element of a the
symmetric matrix

CONFIDENCE LIMITS FOR THE PREDICTED POINTS

- A confidence interval for each predicted value is given by:

$$\hat{y}_i - t \frac{\sigma_{error}}{\nu} \sqrt{[X(X^T X)^{-1} X^T]_{j,j}} < \hat{y}_j < \hat{y}_i + t \frac{\sigma_{error}}{\nu} \sqrt{[X(X^T X)^{-1} X^T]_{j,j}}$$

(9-18)

SUMMARY

- We have seen how fit parameters of a model can be fitted to a data set, using the linear least squares method.
- We found out how to calculate the regression coefficients and how to perform a statistical analysis of the model using ANOVA.
- We also postulated expressions for the confidence limits for the fit parameters and the predicted points